

# Short Course

State Space Models, Generalized Dynamic Systems  
and  
Sequential Monte Carlo Methods,  
and  
their applications  
in Engineering, Bioinformatics and Finance

Rong Chen  
Rutgers University  
Peking University

## Part Three: Advanced Sequential Monte Carlo

### 3.1 Mixture Kalman Filter

#### 3.1.1 Conditional Dynamic Linear Models

#### 3.1.2 Mixture Kalman Filters

#### 3.1.3 Partial Conditional Dynamic Linear Models

#### 3.1.4 Extend Mixture Kalman Filters

#### 3.1.5 Future Directions

### 3.2 Constrained SMC

### 3.3 Parameter Estimation in SMC

### 3.1.1 Conditional Dynamic Linear Models

Indicator  $\Lambda_t$ : (unobserved)

If  $\Lambda_t = \lambda$  then

$$x_t = H_\lambda x_{t-1} + W_\lambda w_t$$

$$y_t = G_\lambda x_t + V_\lambda v_t$$

where  $w_t \sim N(0, I)$  and  $v_t \sim N(0, I)$  and independent.

Given the trajectory of the indicator  $\{\Lambda_1, \dots, \Lambda_t\}$ , the system is linear and Gaussian.

## Example: Tracking a target in clutter

Introducing an indicator  $I_t$  taking values in  $\{0, 1, \dots, m_t\}$ .

$I_t = 0$  true signal missing.  $I_t = i$ , then  $y_t^{(i)}$  is the true signal.

$$\begin{aligned}x_t &= Hx_{t-1} + Ww_t \\y_t^{(i)} &= Gx_t + Vv_t \quad \mathbf{if} \quad I_t = i \\y_t^{(j)} &\sim \mathbf{Unif}(\Delta) \quad \mathbf{if} \quad I_t \neq j\end{aligned}$$

and

$$P(I_t = 0) = p_d \quad \mathbf{and} \quad P(I_t = i) = (1 - p_d)/m_t$$

Given the trajectory of the indicator  $\{I_1, \dots, I_t\}$ , the system is linear and Gaussian.

**Example: Tracking a target with non-Gaussian innovations.**

$$x_t = Hx_{t-1} + Ww_t$$

$$y_t = Gx_t + Vv_t$$

where  $w_t \sim t_{k_1}$ ,  $v_t \sim t_{k_2}$ .

Note that  $t_k = N(0, 1) / \sqrt{\chi_k^2/k}$

Introducing indicators  $\Lambda_t = (\Lambda_{t1}, \Lambda_{t2})$ .

$$\begin{aligned} x_t &= Hx_{t-1} + \frac{\sqrt{k_1}}{\sqrt{\lambda_1}} W w_t & \text{if } \Lambda_{t1} = \lambda_1 \\ y_t &= Gx_t + \frac{\sqrt{k_2}}{\sqrt{\lambda_2}} V v_t & \text{if } \Lambda_{t2} = \lambda_2 \end{aligned}$$

with  $v_t \sim N(0, I)$ ,  $w_t \sim N(0, I)$  and  $\Lambda_{t1} \sim \chi_{k_1}^2$ ,  $\Lambda_{t2} \sim \chi_{k_2}^2$ .

Given the trajectory of the indicator  $\{\Lambda_1, \dots, \Lambda_t\}$ , the system is linear and Gaussian.

**Example: Tracking a target with random (Gaussian) acceleration plus maneuvering**

$$\begin{aligned}x_t &= Hx_{t-1} + Fs_{I_t}u_t + Ww_t \\y_t &= Gx_t + Vv_t\end{aligned}$$

where  $u_t, w_t$  and  $v_t$  are all  $N(0, I)$  independent.

$I_t$  maneuvering status:

$I_t = 0$ , no maneuvering,  $s_0 = 0$

$I_t = 1$ , slow maneuvering,  $s_1$   $I_t = 2$ , fast maneuvering,  $s_2$

With known transition matrix  $P = P(I_{t+1} | I_t)$ .

### 3.1.2 Mixture Kalman Filter:

Let  $\mathbf{y}_t = (y_1, \dots, y_t)$  and  $\Lambda_t = (\Lambda_1, \dots, \Lambda_t)$ .

Note that

$$p(x_t | \mathbf{y}_t) = \int p(x_t | \Lambda_t, \mathbf{y}_t) dF(\Lambda_t | \mathbf{y}_t)$$

and

$$p(x_t | \Lambda_t, \mathbf{y}_t) \sim N(\mu_t(\Lambda_t), \sigma_t^2(\Lambda_t))$$

where

$$KF_t(\Lambda_t) \equiv (\mu_t(\Lambda_t), \sigma_t^2(\Lambda_t))$$

can be obtained from Kalman filter.

$p(x_t | y_1, \dots, y_t)$  is a mixture Gaussian distribution.

**(Sequential) Monte Carlo Filter:**  
**a discrete sample with weight**

$$\{(x_t^{(1)}, w_t^{(1)}), \dots, (x_t^{(m)}, w_t^{(m)})\} \implies p(x_t \mid y_1, \dots, y_t)$$

**Mixture Kalman Filter:**  
**a discrete sample with weight**

$$\{(\boldsymbol{\lambda}_t^{(1)}, w_t^{(1)}), \dots, (\boldsymbol{\lambda}_t^{(m)}, w_t^{(m)})\} \implies p(\boldsymbol{\Lambda} \mid y_1, \dots, y_t)$$

and a random mixture of Normal distributions

$$\sum_{i=1}^m w_t^{(i)} N(\mu_t(\boldsymbol{\lambda}_t^{(i)}), \sigma_t^2(\boldsymbol{\lambda}_t^{(i)})) \implies p(x_t \mid y_1, \dots, y_t).$$

**Hence**

$$E(f(x_t) \mid y_1, \dots, y_t) \approx \sum_{i=1}^m w_t^{(i)} \int f(x) \phi(x; \mu_t(\boldsymbol{\lambda}_t^{(i)}), \sigma_t^2(\boldsymbol{\lambda}_t^{(i)})) dx$$



**Benefit: improved efficiency**

$$\text{Var}[f(x_t) \mid \mathbf{y}_t] \geq \text{Var}[E(f(x_t) \mid \boldsymbol{\Lambda}_t, \mathbf{y}_t) \mid \mathbf{y}_t]$$

**Example:**  $X \sim N(\Lambda, \sigma_1^2)$  and  $\Lambda \sim N(0, \sigma_2^2)$ . **Estimate**  $\mu = E(X)$

**(1) directly sample from**  $X \sim N(0, \sigma_1^2 + \sigma_2^2)$ ,

$$\text{Var}(\hat{\mu}) = \text{Var}\left(\frac{\sum_{i=1}^m X_i}{m}\right) = \frac{\sigma_1^2 + \sigma_2^2}{m}$$

**(2) sample**  $\Lambda \sim N(0, \sigma_2^2)$ .

$$\hat{\mu} = \frac{\sum_{i=1}^m E(X \mid \Lambda_i)}{m} = \frac{\sum_{i=1}^m \Lambda_i}{m}$$

$$\text{Var}(\hat{\mu}) = \text{Var}\left(\frac{\sum_{i=1}^m \Lambda_i}{m}\right) = \frac{\sigma_2^2}{m}$$

**Algorithm:**

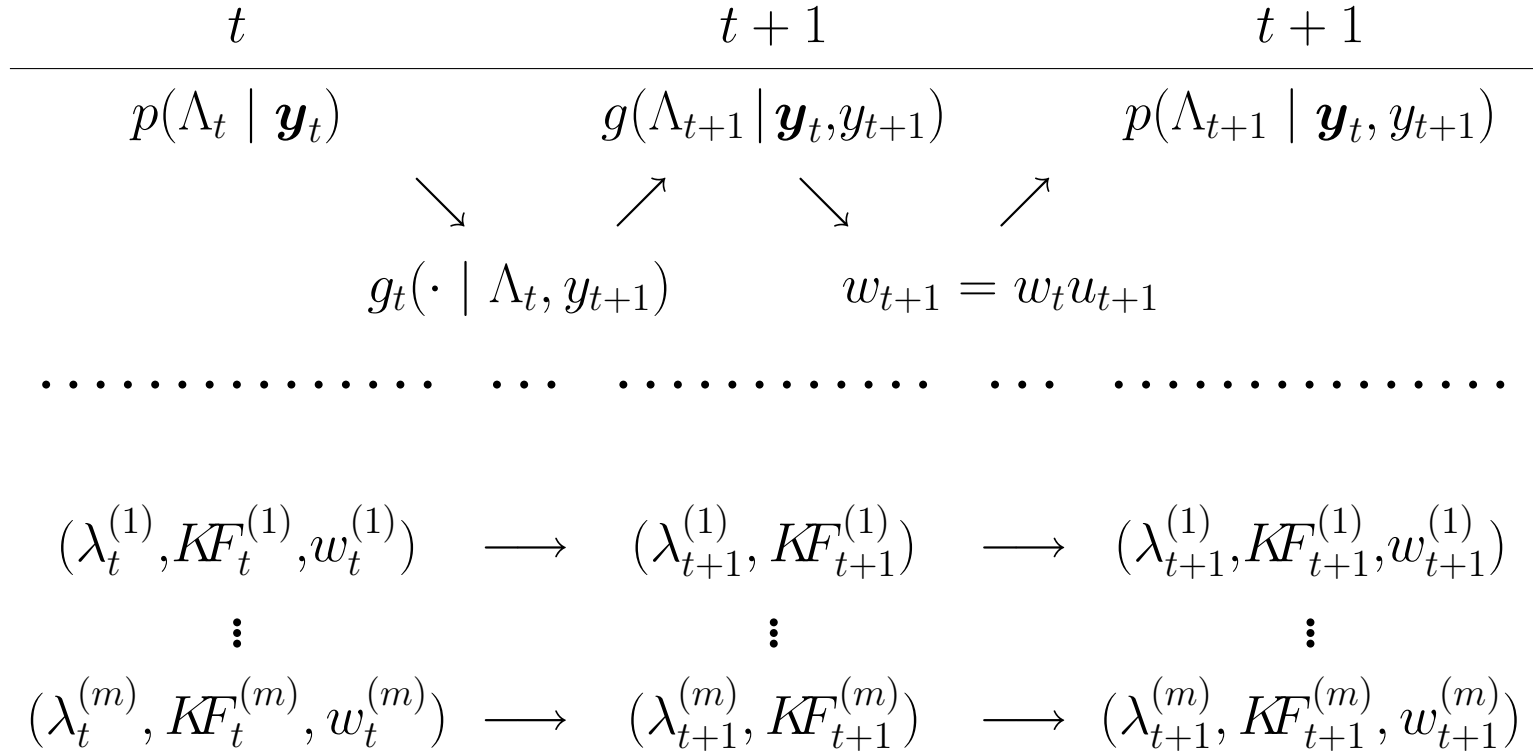
At time  $t$ , we have a sample  $(\boldsymbol{\lambda}_t^{(i)}, KF_t^{(i)}, w_t^{(i)})$

For  $t + 1$ ,

- (1) : generate  $\lambda_{t+1}^{(i)}$  from a trial distribution  $g(\lambda_{t+1} | \boldsymbol{\lambda}_t^{(i)}, KF_t^{(i)}, y_{t+1})$
- (2) : run one step Kalman filter conditioning on  $(\lambda_{t+1}^{(i)}, KF_t^{(i)}, y_{t+1})$  and obtain  $KF_{t+1}^{(i)}$ .
- (3) : calculate the incremental weight

$$u_{t+1}^{(i)} = \frac{p(\boldsymbol{\lambda}_t^{(i)}, \lambda_{t+1}^{(i)} | \mathbf{y}_{t+1})}{p(\boldsymbol{\lambda}_t^{(i)} | \mathbf{y}_t)g(\lambda_{t+1} | \boldsymbol{\lambda}_t^{(i)}, KF_t^{(i)}, y_{t+1})}$$

and the new weight  $w_{t+1}^{(i)} = w_t^{(i)} u_{t+1}^{(i)}$ .



When  $\Lambda_t$  is a discrete r.v. on a finite set, then

(0) : For each  $j = 1, \dots, J$ , run a Kalman filter to obtain

$$u_j^{(i)} = p(y_{t+1} \mid \Lambda_{t+1} = j, KF_t^{(i)})p(\Lambda_{t+1} = j \mid \boldsymbol{\lambda}_t^{(i)})$$

(1) : Sample a  $\lambda_{t+1}^{(i)}$  from the set  $\{1, \dots, J\}$  with probability proportional to  $u_j^{(i)}$ .

i.e. sample a  $\lambda$  from  $p(\Lambda_{t+1} \mid \boldsymbol{\lambda}_t, KF_t, y_{t+1})$

(2) : Let  $KF_{t+1}^{(i)}$  be the one with  $\lambda_{t+1}^{(i)}$ .

(3) : The new weight is

$$w_{t+1}^{(i)} \propto w_t^{(i)} p(y_{t+1} \mid \boldsymbol{\lambda}_t^{(i)}, KF_t^{(i)}) \propto w_t^{(i)} \sum_{j=1}^J u_j^{(i)}$$

When  $\Lambda_t$  is a continuous r.v., a simple (but not optimum) algorithm is

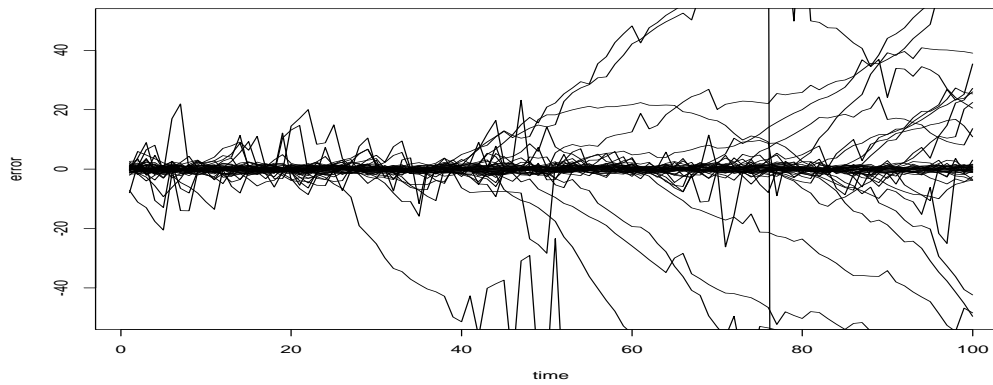
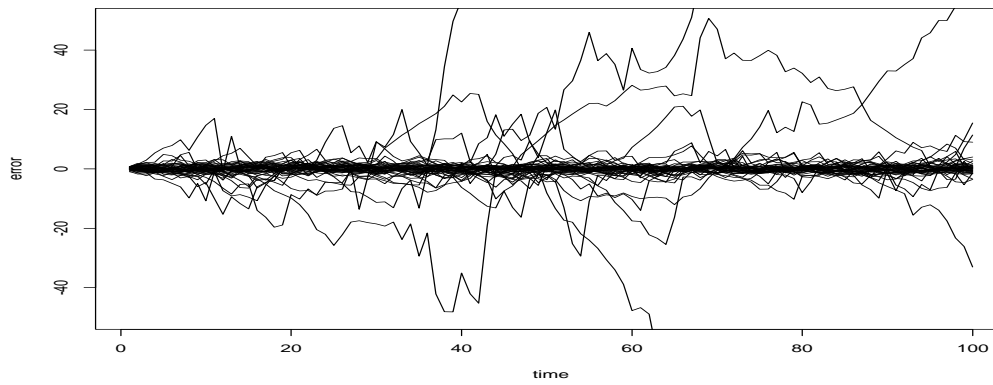
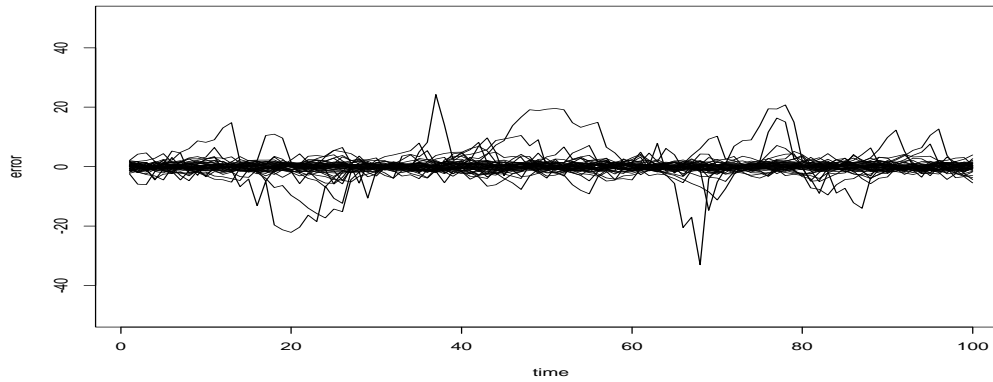
(1) : Sample a  $\lambda_{t+1}^{(i)}$  from  $p(\Lambda_{t+1} \mid \Lambda_t = \lambda_t^{(i)})$

(2) : Run one step Kalman filter conditioning on  $(\lambda_{t+1}^{(i)}, KF_t^{(i)}, y_{t+1})$   
and obtain  $KF_{t+1}^{(i)}$

(3) : The new weight is

$$w_{t+1}^{(i)} = w_t^{(i)} p(y_{t+1} \mid \lambda_{t+1}^{(i)}, KF_t^{(i)})$$

# Tracking in clutter:



## Example: Tracking a target with non-Gaussian innovations

State Equation:

$$\begin{pmatrix} x_t^{(1)} \\ x_t^{(2)} \end{pmatrix} = \begin{pmatrix} 1 & T \\ 0 & 1 \end{pmatrix} \begin{pmatrix} x_{t-1}^{(1)} \\ x_{t-1}^{(2)} \end{pmatrix} + q \begin{pmatrix} T/2 \\ 1 \end{pmatrix} w_t$$

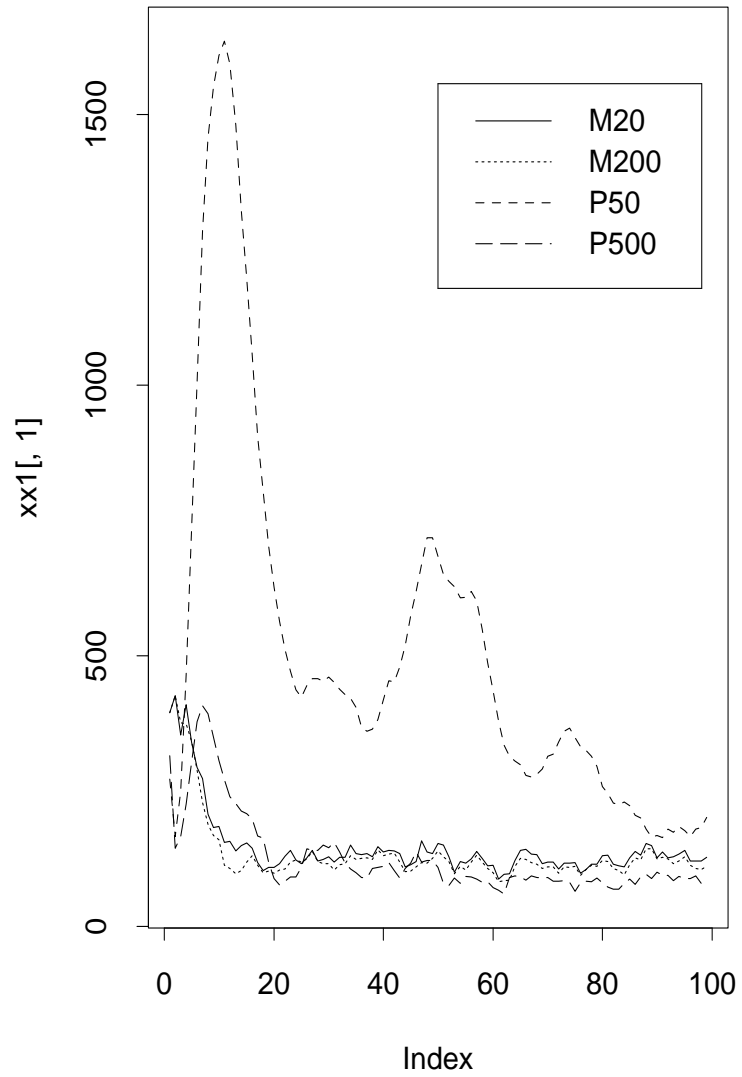
true signal

$$y_t = x_t^{(1)} + r v_t$$

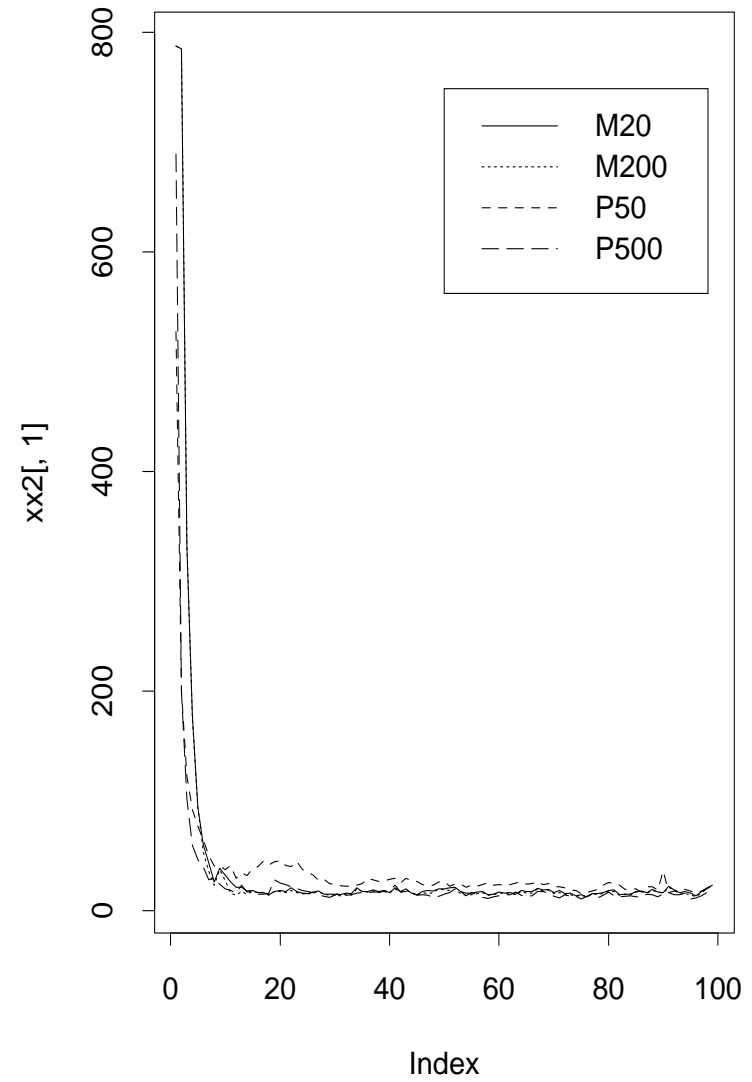
where  $w_t \sim t_3$  and  $v_t \sim t_3$ .

$$T = 1, q^2 = 400/3, r^2 = 1600/3.$$

### MSE x



### MSE speed





noise var	# chains	Particle		MKF	
		cpu time	# miss	cpu time	# miss
$q^2 = 16.0$ $r^2 = 1600$	20	9.49843	72	19.4277	1
	50	20.1622	20	51.6061	1
	200	80.3340	7	181.751	1
	500	273.369	4	500.157	1
	1500	1063.36	3	2184.67	1

### 3.1.3 Partial CDLM

state equation:  $x_t = g_t(x_{t-1}, \varepsilon_t)$

observation equation:  $y_t = h_t(x_t, e_t)$

- Extract the linear and Gaussian components out, and use Kalman Filter (integrating those components out)
- Nonlinear components are dealt with standard Monte Carlo Filters
- Non-Gaussian innovations are dealt with indicators and approximations
- Nonlinear functions are dealt with 'conditional linearization'?

**State:**  $(x_t, x_t^*)$ . **Observations:**  $(y_t, y_t^*)$

$$x_t = g_t(x_{t-1}, x_{t-1}^*, \varepsilon_t) \quad (1)$$

$$x_t^* = H_{x_t} x_{t-1}^* + W_{x_t} w_t \quad (2)$$

$$y_t = h_t(x_t, e_t) \quad (3)$$

$$y_t^* = G_{x_t} x_t^* + V_{x_t} v_t \quad (4)$$

- $x_t, y_t$ : **nonlinear nonGaussian component**
- $x_t^*, y_t^*$ : **conditional linear Gaussian component**
- $H_{x_t}, G_{x_t}, W_{x_t}, V_{x_t}$ : **known matrices given  $x_t$**
- $w_t \sim N(0, I)$  and  $v_t \sim N(0, I)$  and independent.

**Given the trajectory of the NLNG components  $\{x_1, \dots, x_t\}$ , the system (2) (4) is linear and Gaussian.**

## Example: Digital Signal Extraction in Fading Channels

$$\text{State Equations: } \begin{cases} x_t = Hx_{t-1} + w_t \\ \alpha_t = Gx_t \\ s_t \sim p(\cdot | s_{t-1}) \end{cases}$$

$$\text{Observation equation: } y_t = \alpha_t s_t + v_t$$

Or

$$\text{State Equations: } x_t = Hx_{t-1} + w_t$$

$$s_t \sim p(\cdot | s_{t-1})$$

$$\text{Observation equation: } y_t = Gx_t s_t + v_t$$

Example: 2-d target with GPS and IMU sensor.

State:

- position  $p_{1t}, p_{2t}$
- speed  $v_{1t}, v_{2t}$
- (total) acceleration  $a_{1t}, a_{2t}$
- IMU facing  $\theta_t$
- IMU rotational speed  $\psi_t$
- Two motion status:
  - $M_t = 1$ : (roughly) zero acceleration (constant between observations)
  - $M_t = 2$ : (roughly) constant acceleration

$\delta$ : time gap between observations

State equations: ( $M_t = 1$ )

$$p_{it} = p_{it-1} + v_{it-1}\delta + 0.5\delta\varepsilon_{it} \quad i = 1, 2$$

$$v_{it} = v_{it-1} + \varepsilon_{it} \quad i = 1, 2$$

$$a_{it} = \varepsilon_{it} \quad i = 1, 2$$

$$\theta_t = \theta_{t-1} + \psi_{t-1}\delta + 0.5\delta\varepsilon_t^*$$

$$\psi_t = \psi_{t-1} + \varepsilon_t^*$$

$$P(M_t = i \mid M_{t-1} = j) = p_{ij}$$

Similar for  $M_t = 2$

- One can also impose constraints (maps) on the state equations.
- variance of  $\varepsilon_{it}$  depends on platform (walking or vehicle etc)

## Observations:

- $p_{1t}^*, p_{2t}^*$ : (post-processed) GPS signal
- $a_{1t}^*, a_{2t}^*$ : acceleration in the direction of  $\theta_t$
- $\eta_t$ : rotational acceleration

## Observational equations:

$$p_{it}^* = p_{it} + e_{1t} \quad i = 1, 2$$

$$a_{1t}^* = \cos(\theta_t)a_{1t} + \sin(\theta_t)a_{2t} + w_{1t}$$

$$a_{2t}^* = -\sin(\theta_t)a_{1t} + \cos(\theta_t)a_{2t} + w_{2t}$$

$$\eta_t = \psi_t - \psi_{t-1} + w_{3t}$$

Give  $\theta_t, \psi_t, M_t$ , the rest of the system is linear and Gaussian.

Hence,

- $\theta_t, \psi_t, M_t$  are the NLNG stat components
- $\eta_t$  is the NLNG observation component.

## Extended Mixture Kalman Filter:

Let  $\mathbf{y}_t = (y_1, \dots, y_t)$ ,  $\mathbf{y}_t^* = (y_1^*, \dots, y_t^*)$  and  $\mathbf{x}_t = (x_1, \dots, x_t)$ .

Note that

$$\begin{aligned} p(x_t, x_t^* | \mathbf{y}_t, \mathbf{y}_t^*) &= \int p(x_t, x_t^*, \mathbf{x}_{t-1} | \mathbf{y}_t, \mathbf{y}_t^*) d\mathbf{x}_{t-1} \\ &= \int p(x_t^* | \mathbf{x}_t, \mathbf{y}_t^*) p(x_t | \mathbf{x}_{t-1}, y_t, y_t^*) dF(\mathbf{x}_{t-1} | \mathbf{y}_{t-1}, \mathbf{y}_{t-1}^*) \end{aligned}$$

where

$$p(x_t^* | \mathbf{x}_t, \mathbf{y}_t^*) \sim N(\mu_t(\mathbf{x}_t), \sigma_t^2(\mathbf{x}_t))$$

where

$$KF_t(\mathbf{x}_t) \equiv (\mu_t(\mathbf{x}_t), \sigma_t^2(\mathbf{x}_t))$$

can be obtained from Kalman filter.

$p(x_t^* | \mathbf{y}_t, \mathbf{y}_t^*)$  is a mixture Gaussian distribution.



## Inference with EMKF

$$E(f_1(\mathbf{x}_t) \mid \mathbf{y}_t, \mathbf{y}_t^*) \approx \sum_{i=1}^m w_t^{(i)} f_1(\mathbf{x}_t^{(i)})$$

and

$$E(f_2(\mathbf{x}_t^*) \mid \mathbf{y}_t, \mathbf{y}_t^*) \approx \sum_{i=1}^m w_t^{(i)} \int f_2(\mathbf{x}^*) \phi(\mathbf{x}^*; \mu_t(\mathbf{x}_t^{(i)}), \sigma_t^2(\mathbf{x}_t^{(i)})) d\mathbf{x}^*$$

Specially,

$$E(\mathbf{x}_t^* \mid \mathbf{y}_t, \mathbf{y}_t^*) \approx \sum_{i=1}^m w_t^{(i)} \mu_t(\mathbf{x}_t^{(i)})$$

**Benefit: improved efficiency**

$$\text{Var}[f_2(\mathbf{x}_t^*) \mid \mathbf{y}_t, \mathbf{y}_t^*] \geq \text{Var}[E(f_2(\mathbf{x}_t^*) \mid \mathbf{x}_t, \mathbf{y}_t, \mathbf{y}_t^* \mid \mathbf{y}_t, \mathbf{y}_t^*)]$$

### Algorithm:

At time  $t$ , we have a sample  $(\mathbf{x}_t^{(i)}, KF_t^{(i)}, w_t^{(i)})$

For  $t + 1$ ,

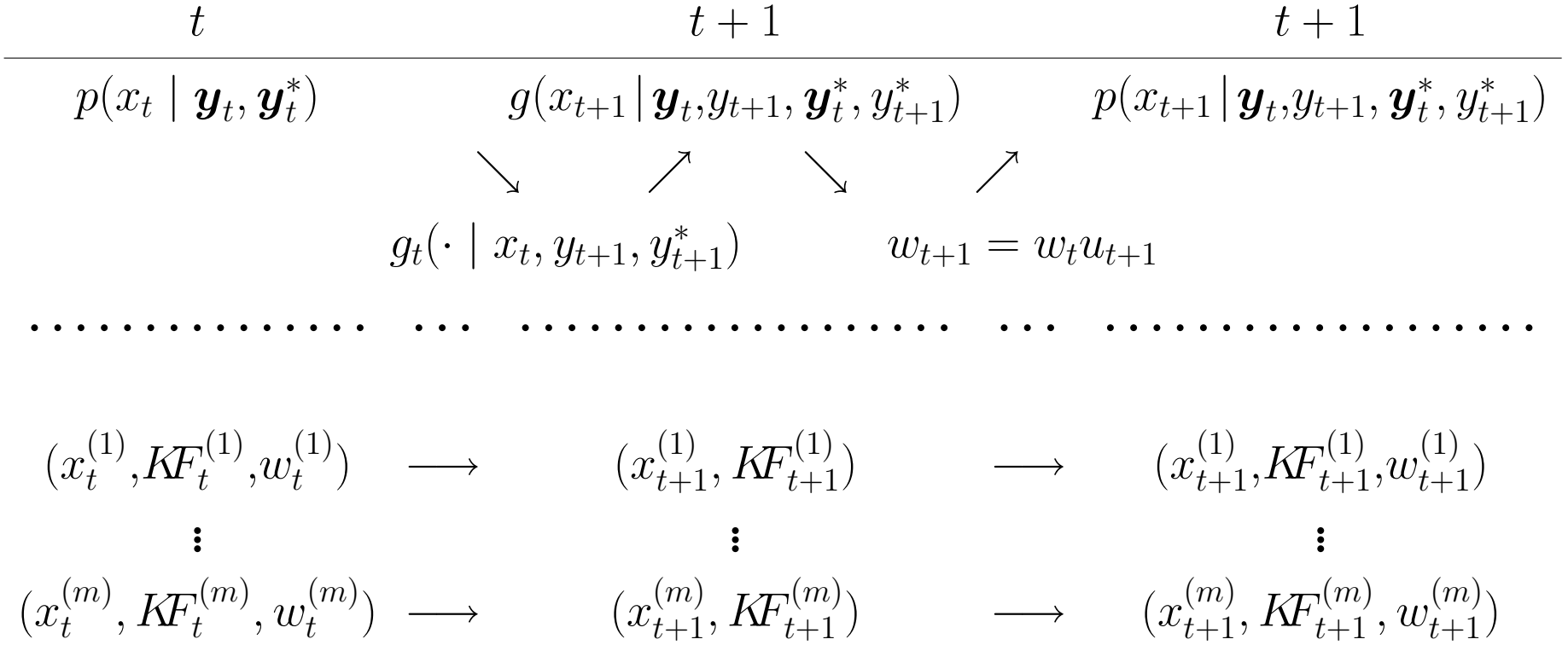
(1) : generate  $x_{t+1}^{(i)}$  from a trial distribution  $g(x_{t+1} | \mathbf{x}_t^{(i)}, KF_t^{(i)}, y_{t+1}, y_{t+1}^*)$

(2) : run one step Kalman filter conditioning on  $(x_{t+1}^{(i)}, KF_t^{(i)}, y_{t+1}^*)$   
and obtain  $KF_{t+1}^{(i)}$ .

(3) : calculate the incremental weight

$$u_{t+1}^{(i)} = \frac{p(\mathbf{x}_t^{(i)}, x_{t+1}^{(i)} | \mathbf{y}_{t+1}, \mathbf{y}_{t+1}^*)}{p(\mathbf{x}_t^{(i)} | \mathbf{y}_t, \mathbf{y}_t^*)g(x_{t+1} | \mathbf{x}_t^{(i)}, KF_t^{(i)}, y_{t+1}^*)}$$

and the new weight  $w_{t+1}^{(i)} = w_t^{(i)} u_{t+1}^{(i)}$ .



**Simple example:**

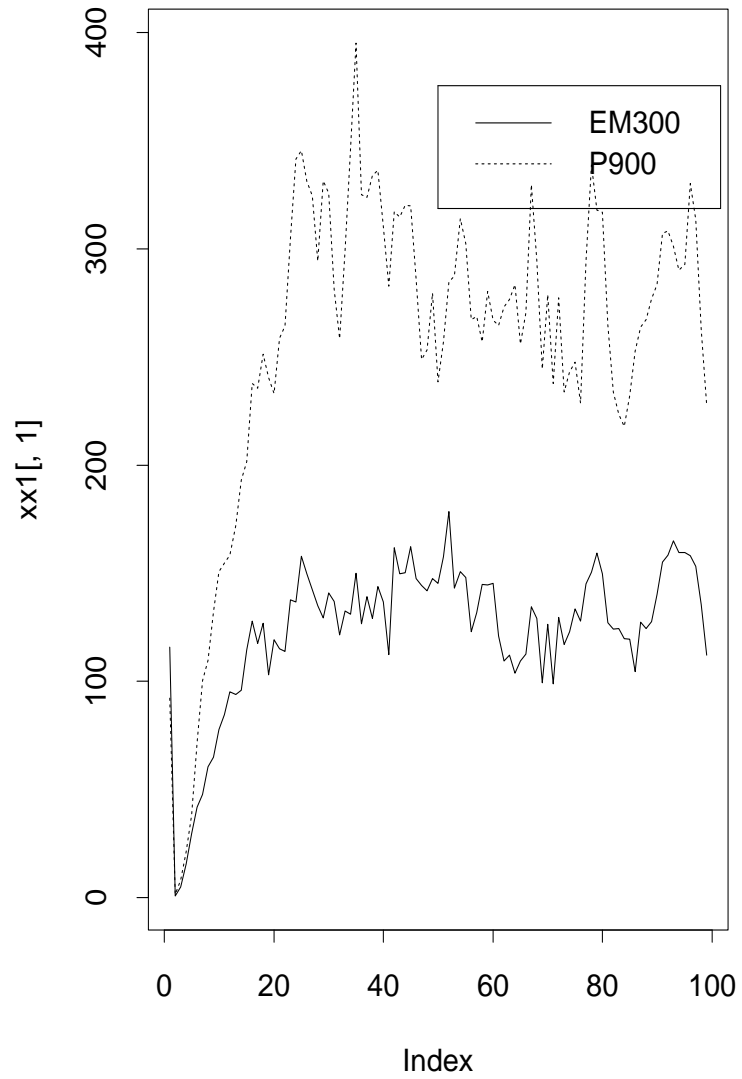
$$x_{t,1} = x_{t-1,2} + 0.2x_{t-1,2}^2 + q_1e_{t1}$$

$$x_{t,2} = x_{t-1,2} + q_2e_{t2}$$

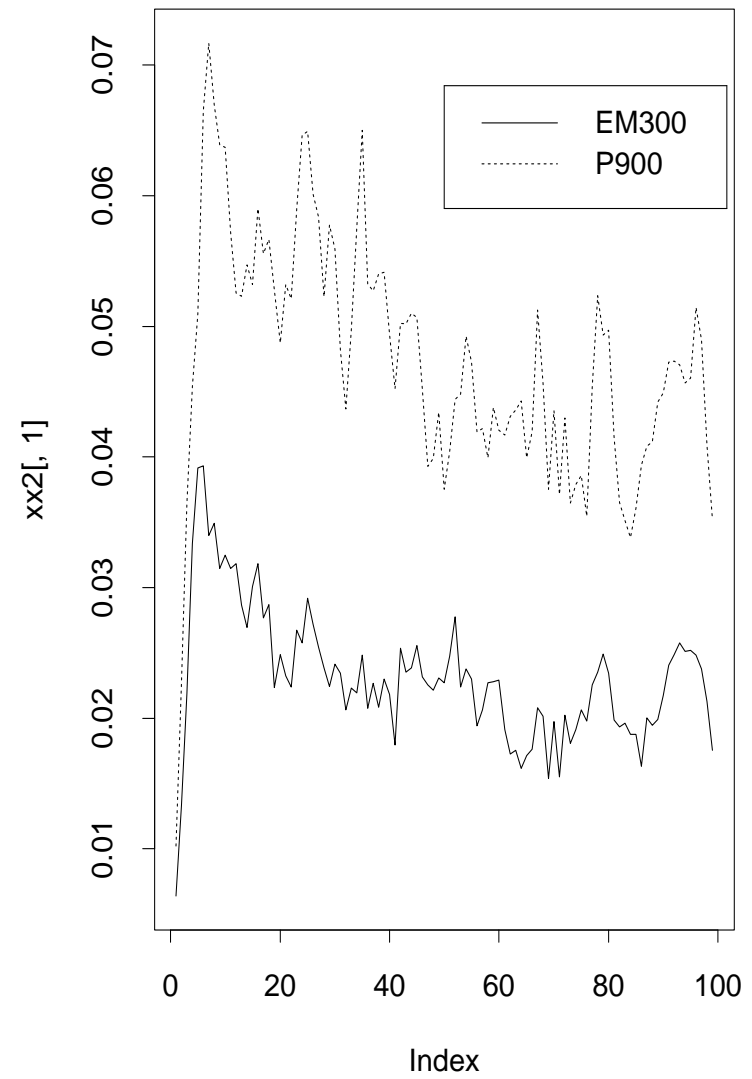
$$y_t = 0.5x_{t,1}x_{t,2} + rv_t$$

**with**  $e_{t1}, e_{t2}, v_t$  **all**  $N(0, 1)$

MSE x1



MSE x2



## Example: Digital Signal Extraction in Fading Channels

$\alpha_t$ : Butterworth filter of order  $r = 3$  i.e. ARMA(3,3)

Cutoff frequency 0.1

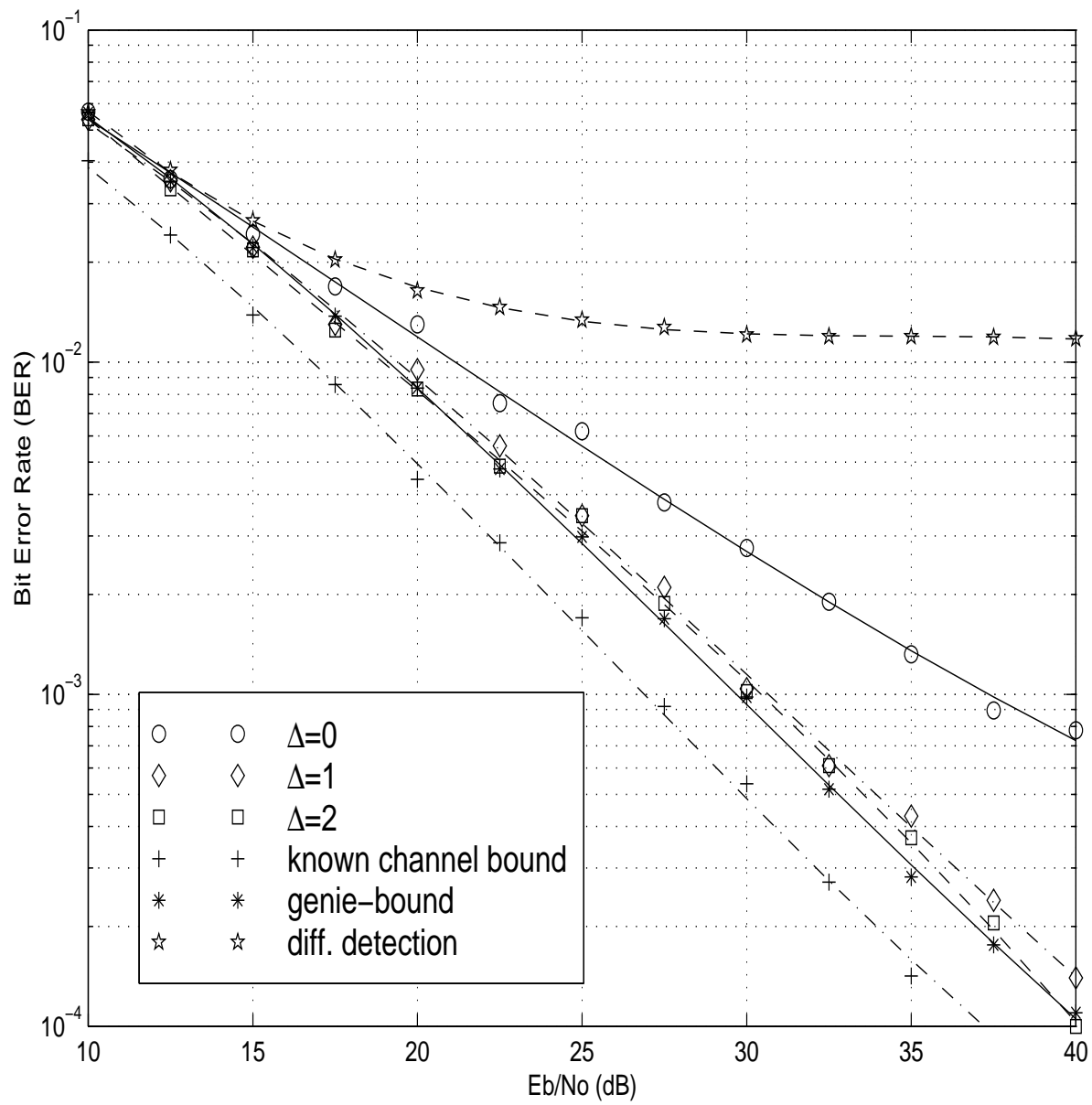
$$s_t = \{-1, 1\}.$$

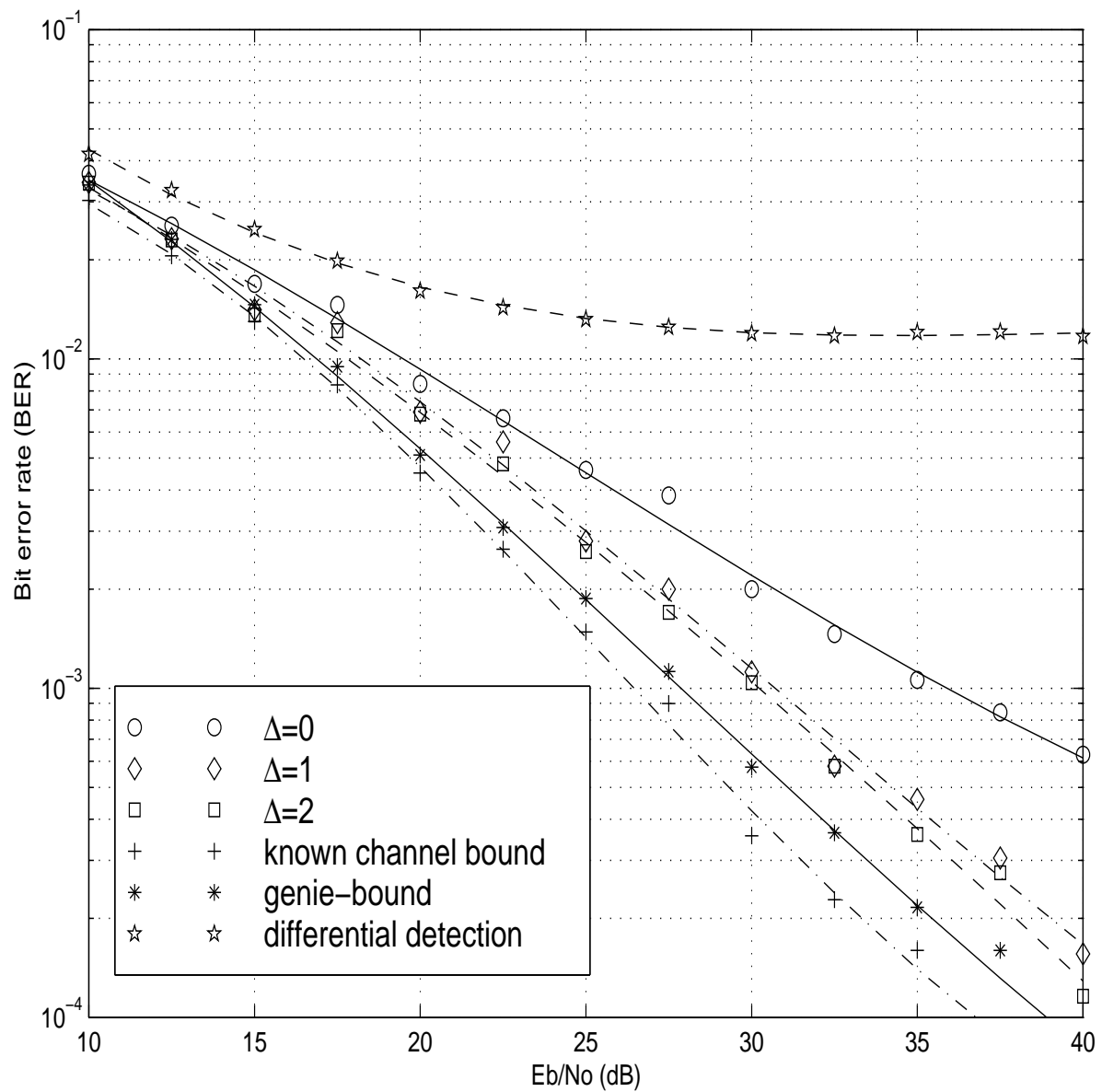
Two cases:

$$v_t \sim N(0, \sigma^2)$$

$$v_t \sim (1 - \alpha)N(0, \sigma_1^2) + \alpha N(0, \sigma_2^2)$$

Extra indicator  $I_t$  for noise.







## What we have done:

- Separate NLNG and Conditional LG components
- Nonlinear components are dealt with standard Monte Carlo Filters
- Use Kalman Filter for the conditional linear and Gaussian component.

## Further Improvement – reduce the number of NLNG components with approximation

- Linear approximation of the nonlinear functions, as EKF.
- Mixture Gaussian approximation of the Non-Gaussian innovations, by introducing indicators
  - conditional on the indicator, innovation is Gaussian
  - $t$  distribution, double exponential, exponential power family, logistic, etc. and the mixture of them!